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To which extent is the membrane potential in a neuron between successive spikes adequately modelled by a (continuous) semimartingale?

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RDS Bielefeld 2010

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between successive spikes

huge litterature modelling the membrane potential between successive spikes as a time homogeneous diffusion, e.g. mean-reverting OU

(overviews in Tuckwell 89 and Lansky-Sato 99, see also Lansky-Sacerdote 01, Ditlevsen-Lansky 05, ...)

or mean reverting CIR

(Lansky-Lanska 87, Giorno-Lansky-Nobile-Ricciardi 88, Lansky-Sacerdote-Tomassetti 95, Ditlevsen-Lansky 06, Brodda-Höpfner 06, ...)

in many data sets, evidence that input terms in the drift are strongly varying with time, and some indication that there might be jumps ...

look to Ito semimartingale models for the membrane potential between successive spikes, and ask questions :

- CIR, OU, ... models, continuous or with jumps, time-homogeneous or not
- model validation relative to shape of diffusion coefficient and drift
- model validation relative to the semimartingale hypothesis

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our data						





data : Kilb and Luhmann, Institute of Physiology, University of Mainz

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example 2 : one pyramidal neuron under different experimental conditions, network activity stimulated by potassium (K) bath, increasing concentration

'Zelle_3': one pyramidal neuron, cortical slice, potassium (K) bath of increasing concentration



data : Luhmann and Kilb, Institute of Physiology, University of Mainz

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CIR type model					

CIR type model $(V_t)_{t\geq 0}$ for the membrane potential between successive spikes :

$$dV_t = (f(t) - V_t) \tau dt + \int y \mu(dt, dy) + \sigma \sqrt{(V_t - K_0)^+} \sqrt{\tau} dW_t$$

with constants $\sigma, \tau > 0$, reference levels $K_0 < K_R < K_E$

- K_E : excitation threshold
- K_R : resting level (:= mean value for m.p. in absence of external stimulus)
- K_0 : lower bound for possible values of the membrane potential

deterministic functions of time

 $f:[0,\infty) \to [K_R,K_E]$, $\widetilde{f}:[0,\infty) \to [0,C]$

modelling external stimulus or degree of activity in the network, and

PRM $\mu(dt, dy)$ on $[0, \infty) \times (0, c]$ with intensity $\tilde{f}(t) \tau dt \nu(dy)$

with measure $u(dy) \,\,\sigma$ -finite on (0,c] such that $\int_{(0,c]} y \,
u(dy) < \infty$

(Poisson random measure μ independent of Brownian motion W; rate of decay $\tau \iff$ 'membrane time constant' $\frac{1}{\tau}$ for biologists)

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time inhomogeneous jump diffusion $(V_t)_t$ as defined above :

proposition 1 : pathwise uniqueness holds, and the m.p. shifted by K_0

$$(V_t - K_0)_{t \geq 0}$$

allows for explicit Laplace transforms

$$\lambda \longrightarrow E\left(e^{-\lambda(V_t-K_0)} \mid (V_s-K_0)=x\right) = \int e^{-\lambda y} P_{s,t}(x,dy)$$

for the transition probabilities, given for fixed x, s < t by

$$\begin{split} \lambda &\longrightarrow \exp\left(-x\Psi_{s,t}(\lambda) - \int_{s}^{t} \left\{ \left[f(\nu) - K_{0}\right]\Psi_{\nu,t}(\lambda) + \widetilde{f}(\nu)\widetilde{\Psi}_{\nu,t}(\lambda) \right\}\tau d\nu \right) \\ \Psi_{\nu,t}(\lambda) &= \frac{e^{-\tau(t-\nu)}\lambda}{1 + \lambda\frac{\sigma^{2}}{2}(1 - e^{-\tau(t-\nu)})} \quad , \quad \widetilde{\Psi}_{\nu,t}(\lambda) = \int \left[1 - e^{-y\Psi_{\nu,t}(\lambda)}\right]\nu(dy) \end{split}$$

in analogy to results of Kawazu-Watanabe TPA 71 for time-homogeneous case (Kawazu-Watanabe 71, Dawson-Li 06, Fu-Li 08, Höpfner 09)

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mp between successive spikes	statistical inference	power variations	varying M	varying <i>p</i>	references	appendix
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<u>remark 1</u>: i) special case $f(\cdot) \equiv f$, $\tilde{f}(\cdot) \equiv \tilde{f}$ constant : then jump diffusion $(V_t - K_0)_{t\geq 0}$ has invariant law with LT

$$\lambda \longrightarrow \exp\left(-\int_{-\infty}^{t} \left\{ \left[\mathbf{f} - \mathbf{K}_{0}\right] \Psi_{v,t}(\lambda) + \widetilde{\mathbf{f}} \, \widetilde{\Psi}_{v,t}(\lambda) \right\} \tau dv \right)$$

independent of t and τ : first term is LT of the Gamma law $\Gamma\left(\frac{2}{\sigma^2}[f-K_0], \frac{2}{\sigma^2}\right)$, second term corresponds to a space-time Poisson mixture of decay processes

('classical' mean reverting CIR models : $f(\cdot) \equiv f$, no jumps $\tilde{f}(\cdot) \equiv 0$)

ii) special case where $f(\cdot)$, $\tilde{f}(\cdot)$ are *T*-periodic functions : have a *T*-periodic semigroup, an invariant probability on the canonical space C[0, T] for *T*-segments in the path of the process, and Harris-recurrence of the Markov chain of *T*-segments : thus we can obtain limit theorems for a large class of functionals of the process $(V_t - K_0)_{t>0}$ (Hopfmer-Kutoyants 10)

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mp between successive spikes	statistical inference	power variations	varying M	varying p	references	appendix
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time homogeneous CIR gives quite good fit for the membrane potential data of example 1 (17Sept08_023) and for the spiking levels 8+9+10 in example 2 (Zelle_3_K_10, K_12, K_15)

among the non-spiking levels 1-7 in example 2, some show good fit to OU type and P type ('bowl shaped diffusion coefficient') diffusions, often with marked time dependence for the input term in the drift

o between successive spikes	statistical inference	power variations	varying M	varying p	conclusion	references	appendix

short remarks on nonparametric statistical inference

restriction to time homogenous model without jumps : assume that a diffusion

$$dX_t := b(X_t) dt + \sigma(X_t) dW_t, \quad t \in [T_0, T_1]$$

is observed on a discrete time grid with suitably small step size $\widetilde{\Delta}$

$$X_{i\widetilde{\Delta}} \ , \ i_0 \leq i \leq i_1 \ , \quad i_0 := \lceil \frac{T_0}{\widetilde{\Delta}} \rceil \ , \ i_1 := \lfloor \frac{T_1}{\widetilde{\Delta}} \rfloor$$

in real data sets : $\widetilde{\Delta} = M\Delta$, Δ the time resolution in the data

we view $a \to \sigma^2(a)$, $a \to b(a)$ as unknown C^1 functions and use kernel estimators with kernel $K(\cdot)$ and bandwidth h > 0 to plot clouds of points

$$(a, \widehat{\sigma^2}(a))_{a \in G}$$
, $(a, \widehat{b}(a))_{a \in G}$, for suitable grids G

in order to make appear a typical shape

kernel $K(\cdot)$: rectangular, triangular, normal ...

mp between successive spikes	statistical inference		varying M	varying p	references	appendix
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using *M*-step Δ -increments in the trajectory for suitable *M*

from H. 07

$$\widehat{\sigma^{2}}(a) := \widehat{\sigma^{2}}_{(\Delta,M,h)}(a) = \frac{\sum_{i=i_{0}}^{i_{1}-M} K\left(\frac{X_{i\Delta}-a}{h}\right) \left(\frac{X_{(i+M)\Delta}-X_{i\Delta}}{\sqrt{\Delta \cdot M}}\right)^{2}}{\sum_{i=i_{0}}^{i_{1}-M} K\left(\frac{X_{i\Delta}-a}{h}\right)}$$
$$\widehat{b}(a) := \widehat{b}_{(\Delta,M,h)}(a) = \frac{\sum_{i=i_{0}}^{i_{1}-M} K\left(\frac{X_{i\Delta}-a}{h}\right) \left(\frac{X_{(i+M)\Delta}-X_{i\Delta}}{\Delta \cdot M}\right)}{\sum_{i=i_{0}}^{i_{1}-M} K\left(\frac{X_{i\Delta}-a}{h}\right)}$$

our choices : bandwith h = 0.01, step multiple M = 20 (check stability under moderate variation of M and h); Δ imposed by structure of data

guideline for estimation of $\sigma^2(\cdot)$: asymptotics $\widetilde{\Delta} \downarrow 0$ as in Florens-Zmirou 93, Hoffmann 99+01, Jacod 10,, for T_1 fixed; for estimation of $b(\cdot)$: assuming ergodicity as in Kutoyants 04, continuous observation up to time T_1 , for $T_1 \rightarrow \infty$

have tightness results in terms of an observable random rate involving

$$\sum_{i=i_0}^{n_1-M} K\left(rac{X_{i\Delta}-a}{h}
ight)$$
 'number of visits near a'

thus 'estimation is reliable at points a where the number of visits is high'

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asymptotics							
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remark on asymptotics (notation : $T_0 = 0$, $T_1 = T$, $i_0 = 0$, $i_1 = \lfloor \frac{T}{\Delta} \rfloor$) : if we were free to vary Δ and T, keeping M fixed, then

proposition : a) under the asymptotics

$$(AS1) T \equiv \mathsf{cst} \ , \ \mathsf{and} \ \Delta \downarrow 0 \ , \ h \downarrow 0 \ , \ h = o\left(\Delta^{1/3}\right)$$

we have tightness of rescaled estimation errors at observable random rate

$$\sqrt{\sum_{i=i_0}^{i_1-M} K\left(\frac{X_{i\Delta}-a}{h}\right)} \left(\widehat{\sigma^2}_{(\Delta,M,h)}(a) - \sigma^2(a)\right)$$

b) under ergodicity/stationarity of $(X_t)_{t\geq 0}$ and under asymptotics

(AS2)
$$\Delta$$
 small , and $T \uparrow \infty$, $h \downarrow 0$, $h = o\left(T^{-1/3}\right)$

we have tightness of rescaled estimation errors

$$\sqrt{\Delta \cdot \sum_{i=i_0}^{i_1-M} K\left(\frac{X_{i\Delta}-a}{h}\right) \left(\widehat{b}_{(\Delta,M,h)}(a)-b(a)\right)}$$

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example							
out of the 10	levels in 'Zel	le_3' (potas	sium bath	n of varyi	ng concen	itration),	
we nick the le	evel 8 which ha	as one isolate	d snike ov		conds of c	hservatio	'n
we plet the t			u spine or				
time ('Zelle	_3_K_10'), and	l apply the at	oove estin	nators to	inter-spik	e-segmen	its
in the membr	ane potential					from H.	07

Zelle_3_K_10.asc



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this example : good fit with a CIR type diffusion

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drift estimation :



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correlation coefficient = -0.877, negative slope = -2.036

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diffusion coefficient estimation :



correlation coefficient = 0.588, positive slope = 0.238

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problems						

similiarly, we can see OU type or P type drift and diffusion cofficient in the non-spiking levels 1–7 in example 2 $\,$

simulated diffusion equivalent : any simulated diffusion path using estimated drift and estimated diffusion coefficient; for our data '17Sept08_023' and for most levels of 'Zelle_3',

pictures as seen above can be reproduced from simulated diffusion equivalents

but we do have a number of problems with 'Zelle_3' :

- some levels might have jumps, in which case we estimate not drift and diffusion coefficient itself, but (drift + associated jump terms) and (diffusion coefficient + associated jump terms)
- strong time inhomogeneities do exist in some levels (e.g. obviously in 1, 2, 4), 'recent past variants' of kernel estimators as above make them appear
- in the spiking level 10 (15 mM of K, 8 spikes) of 'Zelle_3', changing M = 10, 20, 40 sensibly affects estimation of the diffusion cofficient

mp between successive spikes statistical inference power variations varying *M* varying *p* conclusion references appendition occorrection occorrect

<u>model assumption</u>: i) the membrane potential between successive spikes (i.e. on suitable intervals $[T_0, T_1]$ sufficiently away from the spike times) is an Ito semimartingale as in Ait-Sahalia and Jacod AS 2009

$$dX_t = b_t dt + \sigma_t dW_t + \int \kappa \circ \delta(t, y) (\mu - \nu) (dt, dy) + \int \kappa' \circ \delta(t, y) \mu(dt, dy)$$

(processes b, σ , δ , truncation functions $\kappa(y) = y \mathbb{1}_{\{|y| \le 1\}}$, $\kappa'(y) = y \mathbb{1}_{\{|y| > 1\}}$) ii) data : discrete observation of X at time resolution Δ

$$X_{i\Delta} , i_0 \leq i \leq i_1 , \quad i_0 := \lceil \frac{T_0}{\Delta} \rceil , i_1 := \lfloor \frac{T_1}{\Delta} \rfloor$$

aim : apply 'test for jumps' in Ait-Sahalia and Jacod AS 2009 to data and ask

• can the data be viewed as a continuous semimartingale?

• can the data be viewed as a semimartingale having jumps?

and finally -the test by Ait-Sahalia and Jacod establishing a dichotomy- ask

• can we assume a semimartingale model for our membrane potential data? there will be a surprising answer ...

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consider $T_0 = 0$, $T_1 = T$, and also $0 \le S < T$; on]]S, T]]:

if we were free to work with arbitrary $\widetilde{\Delta}\downarrow 0$, we would use usual *p*-variations

$$\widehat{B}_{S,T}(p,\widetilde{\Delta}) := \sum_{i=1}^{\lfloor (T-S)/\widetilde{\Delta} \rfloor} \left| X_{S+i\widetilde{\Delta}} - X_{S+(i-1)\widetilde{\Delta}} \right|^p \quad \text{as} \quad \widetilde{\Delta} \downarrow 0$$

defined w.r.t. non-intersecting intervals (work of Jacod, Barndorff-Nielsen, ...); in order to isolate Gaussian parts as $\widetilde{\Delta} \downarrow 0$, we would use truncated *p*-variations

$$\check{B}^{\epsilon}_{S,T}(p,\widetilde{\Delta}) := \sum_{i=1}^{\lfloor (T-S)/\widetilde{\Delta} \rfloor} \mathbb{1}_{\left\{ \left| X_{S+i\widetilde{\Delta}} - X_{S+(i-1)\widetilde{\Delta}} \right| < \left[\widetilde{\Delta} \right]^{\frac{1-\epsilon}{2}} \right\}} \left| X_{S+i\widetilde{\Delta}} - X_{S+(i-1)\widetilde{\Delta}} \right|^{p}$$

motivated by LIL (Mancini SJS '09, SPA '10, Podolskii and Ziggel SISP '10)

in our data, it is impossible to make work convergence results of this type ... can only mimick $\widetilde{\Delta} \downarrow 0$ in a very poor way by considering

$$\Delta := M\Delta \quad , \quad M = \dots, 5, 4, 3, 2, 1$$

decreasing multiples of the step size Δ prescribed by the data

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for $T_0 = 0$, $T_1 = T$ and step size Δ in the data, write $i_0 = 0$, $i_1 = \lfloor \frac{T}{\Delta} \rfloor$

for varying *M* and *p*, we will use <u>M-step Δ -increments</u> in the trajectory, and will mix *p*-variations as above

$$V_{0,T,\infty}(\rho,\Delta,M) := \frac{1}{M} \sum_{\ell=0}^{M-1} \widehat{B}_{\ell\Delta,T}(\rho,M\Delta) = \frac{1}{M} \sum_{i=i_0}^{i_1-M} |X_{(i+M)\Delta} - X_{i\Delta}|^{\rho}$$

and work with overlapping intervals; the M summands in the second term being 'almost identical', we identify (heuristics!)

$$V_{0, \mathcal{T}, \infty}(p, \Delta, M) \approx \widehat{B}_{0, \mathcal{T}}(p, M\Delta)$$

then transfer established convergence results for $\widehat{B}_{0,\tau}(p,\widetilde{\Delta})$ as $\widetilde{\Delta} \downarrow 0$ to $V_{0,\tau,\infty}(p,\Delta,M)$ for decreasing $M = \ldots, 3, 2, 1$

our truncated *p*-variations will be

$$V_{0,T,Z}(p,\Delta,M) := \frac{1}{M} \sum_{i=i_0}^{i_1-M} \mathbf{1}_{\left\{ \left| X_{(i+M)\Delta} - X_{i\Delta} \right| < 3 \cdot [M\Delta]^{\frac{1}{2}} \cdot Z \right\}} \left| X_{(i+M)\Delta} - X_{i\Delta} \right|^p$$

where we play around with the deterministic truncation factor Z, $Z \uparrow \infty$ in spiking neurons : several disjoint diffusion-like segments : we add up the corresponding terms coming from the segments

(Höpfner 10)

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varying multiples M of Δ when p = 2, 4 is fixed

on the basis of the above heuristics consider 2- and 4-variations :

4-variations : the test for jumps in Ait-Sahalia and Jacod AS 2009

 $\begin{cases} X \text{ with jumps : } \widehat{B}_{0,T}(4,\widetilde{\Delta}) & \xrightarrow{\Delta \downarrow 0} & \sum_{0 < s \le T} |X_s - X_{s^-}|^4 \quad (\text{str. pos. limit}) \\ X \text{ continous : } \widehat{B}_{0,T}(4,\widetilde{\Delta}) & \xrightarrow{\widetilde{\Delta} \downarrow 0} & \widetilde{\Delta} \cdot \left[m_4 \int_0^T \sigma_s^4 ds \right] \quad (\text{linearity in } \widetilde{\Delta}) \end{cases}$ should be rephrased in our setting as follows : while decreasing $M = \ldots, 3, 2, 1$, $\begin{cases} X \text{ with jumps : } V_{0,T,\infty}(4,\Delta,M) \text{ stabilizes at a strictly positive 'limit'} \\ X \text{ continous : } V_{0,T,\infty}(4,\Delta,M) \text{ is approx. linear in } M, \text{ and null at 0} \end{cases}$

2-variations : the well known convergence

$$\widehat{B}_{0,T}(2,\widetilde{\Delta}) \quad \stackrel{\widetilde{\Delta}\downarrow 0}{\longrightarrow} \quad [X,X]_T = \int_0^T \sigma_s^2 ds + \sum_{0 < s \leq T} |X_s - X_{s^-}|^2$$

should translate as follows : for decreasing and sufficiently small values of M,

 $V_{0,T,\infty}(2,\Delta,M)$ is strictly positive and flat in M ▲□▶ ▲□▶ ▲目▶ ▲目▶ 目 - シ۹ペ

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non-spiking neurons						

non- or very rarely spiking neurons : 4-variations as a function of M

data 'Zelle_3_K_10' (level 8), one isolated spike over 60 seconds :



Zelle_3_K_10.asc, whole trajectory, 4-variations depending on M

good linearity for all $Z \in \{1, 2, 4, 8, 10, 16, 32, \ldots\}$, no changes above $Z \ge 8$ periodic deformations (circuits/loops in the neuronal slice?) no clear hint to presence of jumps : behaviour of a (continuous) semimartingale similiar pictures for all non-spiking levels 1–7 of 'Zelle_3'

mp between successive spikes	statistical inference		varying M	varying p	references	appendix
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non- or very rarely spiking neurons : 2-variations as a function of M same data 'Zelle_3.K_10', level 8 , one isolated spike :



Zelle_3_K_10.asc, whole trajectory, 2-variations depending on M

essentially flat in M for $M \ge 15$, for all values of truncation factor Zsame periodic deformations observable as above (circuits ?) 'microstructure noise' visible for small M, probably due the recording electrode (similiar in all levels of 'Zelle_3', not present in '17Sept08_023')

behaviour of a semimartingale, up to noisy observation similiar pictures for all non-spiking levels 1–7 of 'Zelle_3'

mp between successive spikes	statistical inference	power variations	varying M	varying p	references	appendix
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spiking neurons						

sufficiently frequently spiking neurons : 4-variations as a function of M

in a semimartingale, for p = 4 fixed and M = ..., 3, 2, 1 decreasing, 4-variations should be either linear or converging to a strictly positive limit

inspecting the frequently spiking neuron '17Sept08_023' (\approx 50 spikes over 60 seconds, no 'noise'), we see none of both : semimartingale property violated



(truncation factors $Z \in \{1, 2, 4, 8, 10, 16, 32, 64, \ldots\}$, no changes for $Z \ge 32$)

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mp between successive spikes	statistical inference	power variations	varying M	varying p	references	appendix
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sufficiently frequently spiking neurons : 2-variations as a function of M

in a semimartingale, for p = 2 fixed and $M = \ldots, 3, 2, 1$ decreasing, 2-variations should be flat in M

for the spiking level 10 ('Zelle_3_K_15', 8 spikes over 60 seconds), we see



Zelle_3_K_15.asc, 60 sec, 2-variations depending on M

M (using increments over M * 6e-04 sec, with 6e-04 the time resolution in data set)

(truncation factors as above, noisy observation visible in small *M*-values)

semimartingale property violated : similiar pictures also for level 9 of 'Zelle_3' (18 spikes), and for '17Sept08_023'

varying p when the multiple M of Δ is fixed

the test for jumps in Ait-Sahalia and Jacod AS 09

for
$$\xi$$
 with jumps : $\widehat{B}_{0,T}(p, 2\widetilde{\Delta}) \approx \begin{cases} \widehat{B}_{0,T}(p, \widetilde{\Delta}) & \text{for } 2 \leq p < \infty \\ 2^{\frac{p}{2}-1} \widehat{B}_{0,T}(p, \widetilde{\Delta}) & \text{for } 0 < p < 2 \end{cases}$
for ξ continuous : $\widehat{B}_{0,T}(p, 2\widetilde{\Delta}) \approx 2^{\frac{p}{2}-1} \widehat{B}_{0,T}(p, \widetilde{\Delta}) & \text{for } 0$

(for $\widetilde{\Delta}$ small enough) translates in our setting as follows : empirical log-ratios

$$p \longrightarrow \log \frac{V_{0,T,\infty}(p,\Delta,2M)}{V_{0,T,\infty}(p,\Delta,M)} \approx \log \frac{\widehat{B}_{0,T}(p,2M\Delta)}{\widehat{B}_{0,T}(p,M\Delta)}$$

(heuristics as above !) should be close to the deterministic function

$$\left\{ \begin{array}{ll} p \to \min\left\{ \left(\frac{p}{2} - 1\right)\log 2 \,,\, 0\right\} & \text{ on } 0$$

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in our data, we work again with M = 20

mp between successive spikes	statistical inference	power variations	varying M	varying p	references	appendix
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non-spiking neurons						

non- or very rarely spiking neurons : empirical log-ratios as a function of p

data 'Zelle_3_K_10', level 8, one isolated spike :



Zelle_3_K_10.asc, 60 sec, test for jumps through p-variations functional in p

(truncation factors $Z \in \{1, 4, 16, 64, 256\}$, no changes above $Z \ge 16$)

essentially linear in p: behaviour of a continuous semimartingale up to two exceptions, the non-spiking levels 1–7 produce this type of picture

mp between successive spikes	statistical inference	varying M	varying p	references	appendix
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among the non-spiking levels 1–7 of 'Zelle_3', the present methods detects jumps in some cases : here level 4 ('Zelle_3_K_6'), truncation fact. as above :

Zelle_3_K_6.asc, 60 sec, test for jumps through p-variations functional in p



behaviour of a semimartingale with jumps

<u>so far : can conclude from our data :</u> in non-spiking or very rarely spiking regimes, the membrane potential is adequately modelled as a semimartingale, in some cases with jumps

mp between successive spikes	statistical inference	power variations	varying M	varying p	references	appendix
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spiking neurons						

sufficiently frequently spiking neurons : empirical log-ratios as a function of p

in a semimartingale, for M fixed and varying p, empirical log-ratios should be either linear in p (with known slope), or linear with truncation at 0 for $p \ge 2$; in 'Zelle_3_K_12', level 9, 18 spikes, none of both :



Zelle_3_K_12.asc, 60 sec, test for jumps through p-variations functional in p

(truncation factors $Z \in \{1, 4, 16, 64, 256\}$, no changes above $Z \ge 16$)

semimartingale property violated; similiar pictures for level 10 of 'Zelle_3' (8 spikes), and for '17Sept08_023'

the disturbing conclusion									

frequently spiking neuron '17Sept08_023', spiking levels 9+10 of 'Zelle_3': power variations calculated from data do <u>not</u> fit into the dichotomy established by Jacod's test for jumps for Ito semimartingales :

hence the membr. pot. between successive jumps in <u>frequently spiking neurons</u> IS NOT ADEQUATELY MODELLED BY A SEMIMARTINGALE ...

this forbids any semimartingale based approach to 'information transmission' in neurons ... and destroys a large number of papers analyzing interspike times in terms of level crossing times of diffusions

whereas in the <u>non-spiking regimes</u> of levels 1-7 of 'Zelle_3', and even in level 8 where one isolated spike occurs, our data (up to effects of 'noise')

VALIDATE THE SEMIMARTINGALE HYPOTHESIS.

Hence there is a strong need for neuron models which allow to understand non-spiking and spiking regimes simultaneously ...

mp between successive spikes	statistical inference	power variations	varying M	varying p	conclusion	references	appendix

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mp between successive spikes	statistical inference	power variations	varying M	varying <i>p</i>	conclusion	references	appendix
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mp between successive spikes	statistical inference		varying M	varying p	references	appendix
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appendix



spikes are generated when the membran potential V_t in the soma is high enough

mp between successive spikes	statistical inference	power variations	varying <i>M</i>	varying <i>p</i>	references	appendix
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heuristics : why should the membrane potential between successive spikes be modelled as a (jump) diffusion process?

one neuron has $O(10^4)$ synapses, pprox 90% excitatory, pprox 10% inhibitory



contribution of incoming spikes to the membrane potential via : exciting synapsis (L), inhibitory synapsis (M), exciting and inhibitory combined (R)

synapses \hookrightarrow dendrites \hookrightarrow soma : additivity and exponential decay